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Yield Model Development

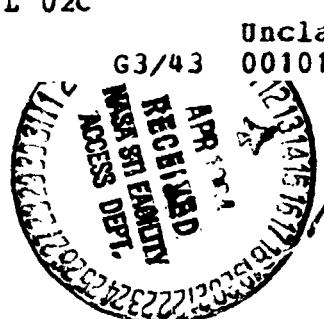
ARGENTINA WHEAT YIELD MODEL

SUSAN L. CALLIS
CLARENCE SAKAMOTO

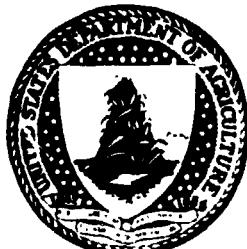
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16. Abstract Five models based on multiple regression were developed to estimate wheat yields for the five wheat-growing provinces of Argentina. Meteorological data sets were obtained for each province by averaging data for stations within each province. Predictor variables for the models were derived from monthly total precipitation, average monthly mean temperature, and average monthly maximum temperature. Buenos Aires was the only province for which a "trend variable" was included because of increasing trend in yield due to technology from 1950 to 1963.			
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ARGENTINA WHEAT YIELD MODELS

by

Susan L. Callis

and

Clarence Sakamoto

AISC Models Branch

January 16, 1984

INTRODUCTION

The purpose of this study was to select monthly weather variables that could be used in models to predict wheat yields for the five main wheat-growing provinces of Argentina: Buenos Aires, Cordoba, Entre Rios, La Pampa and Santa Fe (see Figure 1). Each province was treated as a separate entity.

Most of Argentina's wheat-growing area is located in a humid subtropical climate known as the Pampa Humida. High temperatures in the winter can have an adverse effect on the wheat yield in the northern sections. Excessive rainfall in Entre Rios, where annual rainfall ranges from 900 to 1100 mm, can also be a problem during the growing season. The western edge of the wheat area, however, is semi-arid with warm to hot summers. There, drought and high temperatures can be a problem during the growing season.

Wheat is planted from early May through July. Harvest is generally from November through January.

METHOD

Three indices representing available soil moisture, monthly precipitation, and monthly maximum temperature (Sakamoto, 1976) were used in multiple regression models. They include: monthly Z-index, ET (evapotranspiration) minus ET (climatically appropriate evapotranspiration), and precipitation minus PET (potential evapotranspiration). Terms are defined in the Appendix. Large positive Z-index, P-PET, and ET-ET values suggest wet conditions.

The regression equation is:

$$\hat{Y} = \alpha + B_1 T + B_2 T X_i + B_3 R_i + B_4 Z_i + B_5 (P-PET)_i + B_6 (ET-ET)_i + E$$

where

\hat{Y} = Estimated yield,

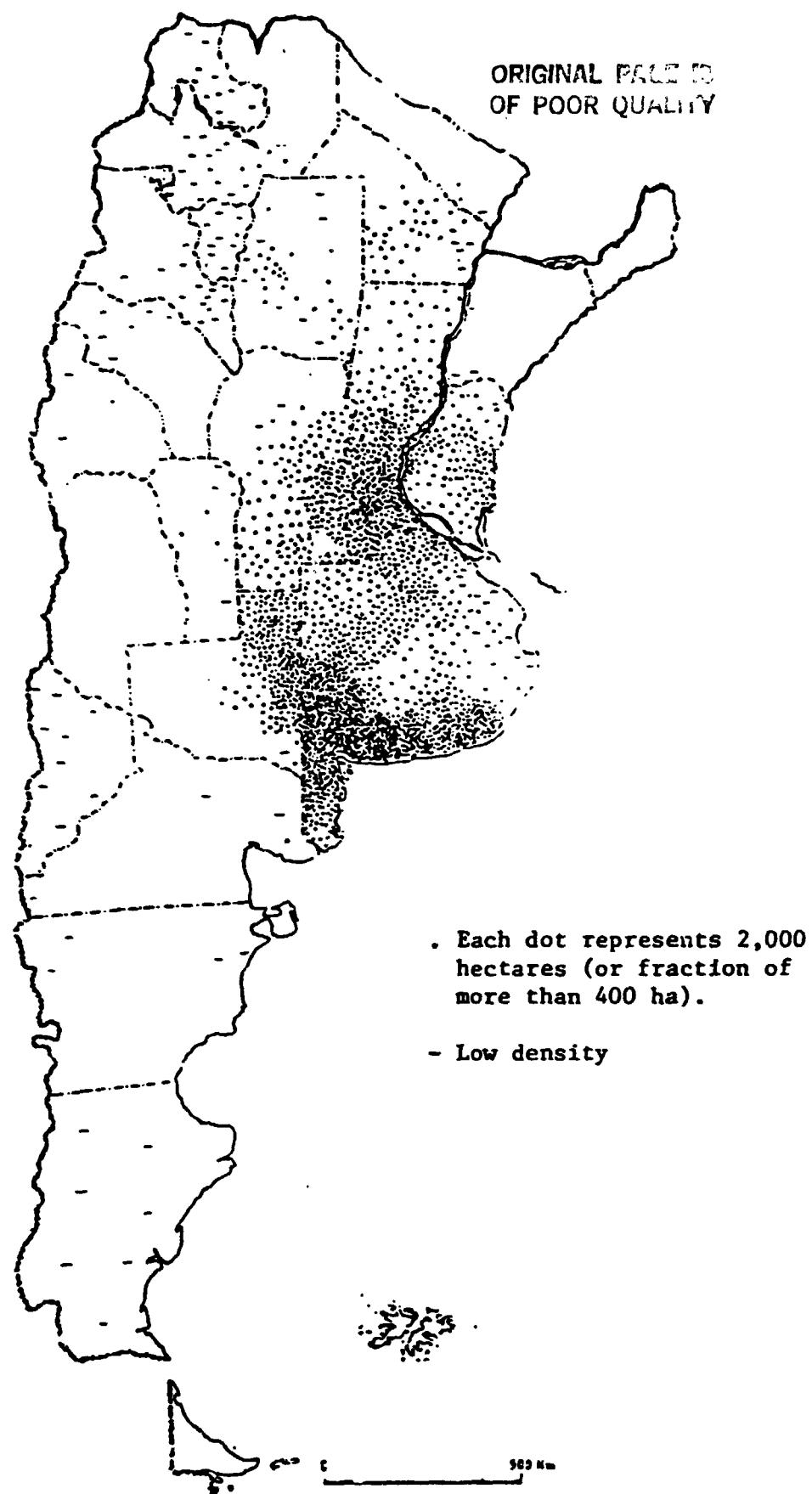


Figure 1. Density of area sown to wheat in Argentina, 1971-72.
(Total area in wheat: 4,986,000 ha)

α = Constant,
 B_j = Coefficients of the variables $j = 1-6$,
 T = Trend,
 TX_j = Maximum temperature for month j ,
 R_j = Total precipitation for month j ,
 Z_j = Z-index for month j ,
 $(P-PET)_j$ = Precipitation minus PET for month j ,
 $(ET-\hat{ET})_j$ = ET minus \hat{ET} for month j , where $\hat{ET} = K \cdot PET$ and $K = \bar{ET}/\bar{PET}$,
and
 E = Unexplained error.

Separate data sets were used for each province. Variables tried and selected for each model differed according to the climate and agricultural practices of each province. Buenos Aires was the only province for which a trend variable was included.

In developing the models, various procedures of the Statistical Analysis System (SAS Institute Inc., 1979) were used. The procedures used and the operations performed with each are listed in the Appendix. Combinations of variables with the highest R^2 which included variables significant at (or close to) the 10 percent level and those agronomically meaningful were chosen for the models.

DATA

The Argentina crop data were obtained through the Latin American Branch of the Economic Research Service of the United States Department of Agriculture (M. Mielke, personal communication, 1980. The crop data set used was set up with the year of yield as the year of harvest. The growing season in Argentina may extend into January of the following year before harvest begins. Our interest is in the year of planting and when the weather impacts the crop, which in this case would be year-1.

Meteorological data were prepared using several different sources, including Monthly Climatic Data for the World and the Servicio Meteorológico Nacional in Argentina (R.E. Jensen, C.M. Sakamoto and S.E. Mummert; August 1974). The years between 1950 and 1970 were used to model since the greatest number of stations had the most complete meteorological data for these years. From a general meteorological data file of Argentina stations, separate data sets were created for each of the five wheat provinces. Table 1 lists the stations and weights used to derive each meteorological data set. Groups of stations were weighted according to the contribution of their area to the country's wheat production. Figure 2 shows the location of each station.

YIELD MODELS

Buenos Aires

A plot of yield vs. year for Buenos Aires reveals an increasing trend in technology from 1950 to about 1963 (see Figure 3). Therefore, a "trend variable" was chosen for this period.

The model for Buenos Aires was one of the most difficult to define, probably because this province covers such a large territory. Plots of possible weather variables versus detrended yield did not indicate any significant linear relationships. Variables that one would be inclined to choose on the basis of crop calendar and critical weather during growing season did not show strong correlations with yield.

After trying many combinations of variables in regressions, the following model was selected:

Linear Trend 1950-1963
ETMETH4.....April ET minus \hat{ET}
R7.....July total precipitation
ZINDEX8.....August Z-index
TX9.....September average maximum temperature.

The signs of the coefficients of the model seem reasonable. The negative coefficient in July indicates too much rain during planting has a negative

BUENOS AIRES

Pergamino
Junin
San Miguel
Buenos Aires
Nueve de Julio
Trenque Lauquen
Las Flores

Averaged and weighted 25%

Patagones
Fortin Mercedes
Tres Arroyas
Trenque Lauquen
Macachin

Averaged and weighted 55%

Azul
Tres Arroyas
Balcarce

Averaged and weighted 20%

CORDOBA

Pilar
Bell Ville
Rio Cuarto
Trenque Lauquen

Averaged

ENTRE RIOS

La Paz
Concordia
Parana
Las Delicias
Victoria

Averaged

LA PAMPA

Patagones
Santa Rosa
Macachin
General Acha
Fortin Mercedes
Victorica
Trenque Lauquen

Averaged

SANTA FE

Ceres
Esperanza
Angel Gallardo
Bell Ville

Averaged and weighted 23%

Bell Ville
Rosario
Casilda
Pergamino

Averaged and weighted 77%

Table 1. Meteorological Stations and Weights Used to Develop Regression Models for Argentina Wheat

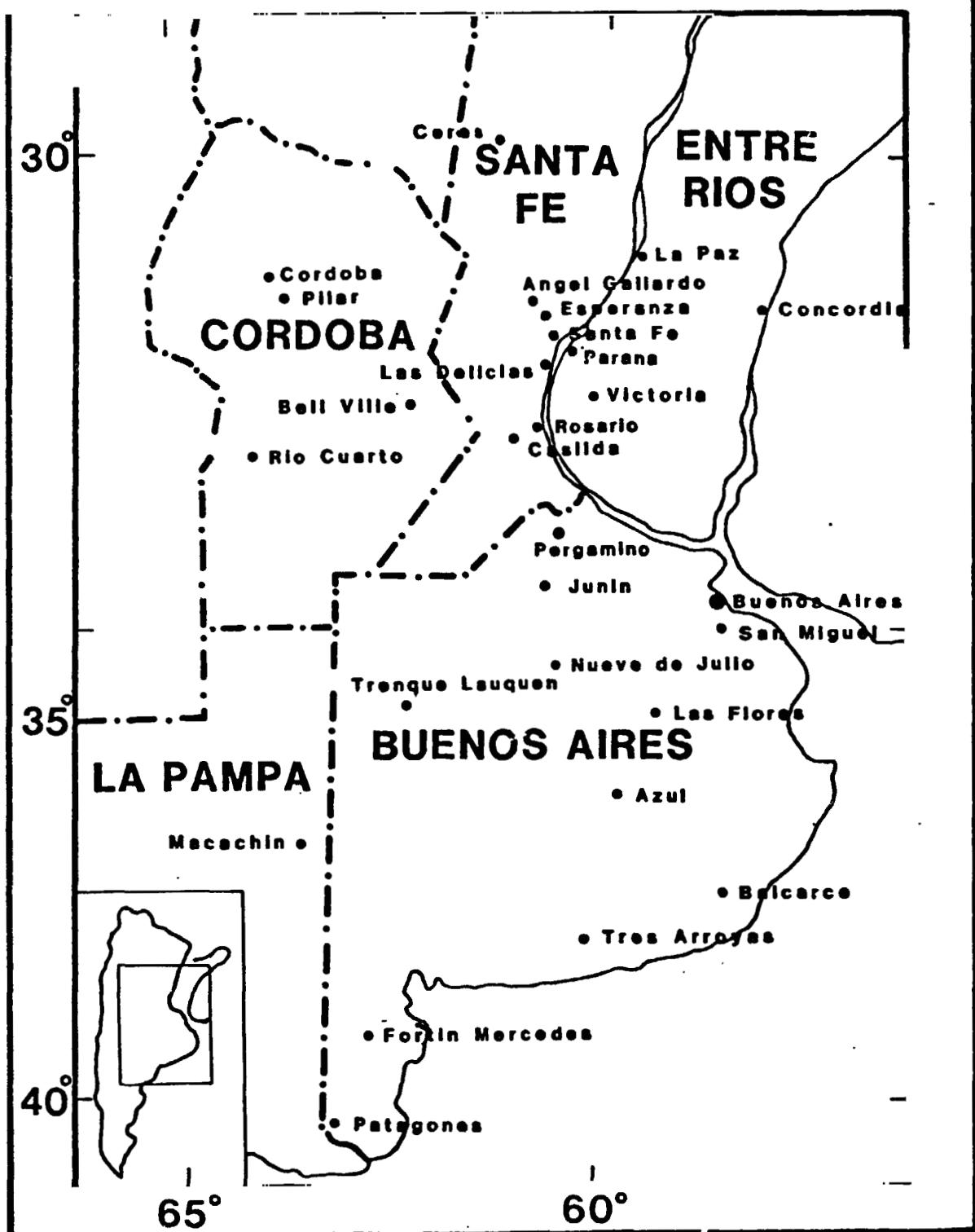


Figure 2. Five major agricultural provinces in Argentina and locations of meteorological stations used in model development.

STATISTICAL ANALYSIS SYSTEM
PLOT OF YIELD*YEAR

LEGEND: A = 1 OBS, B = 2 OBS, ETC.

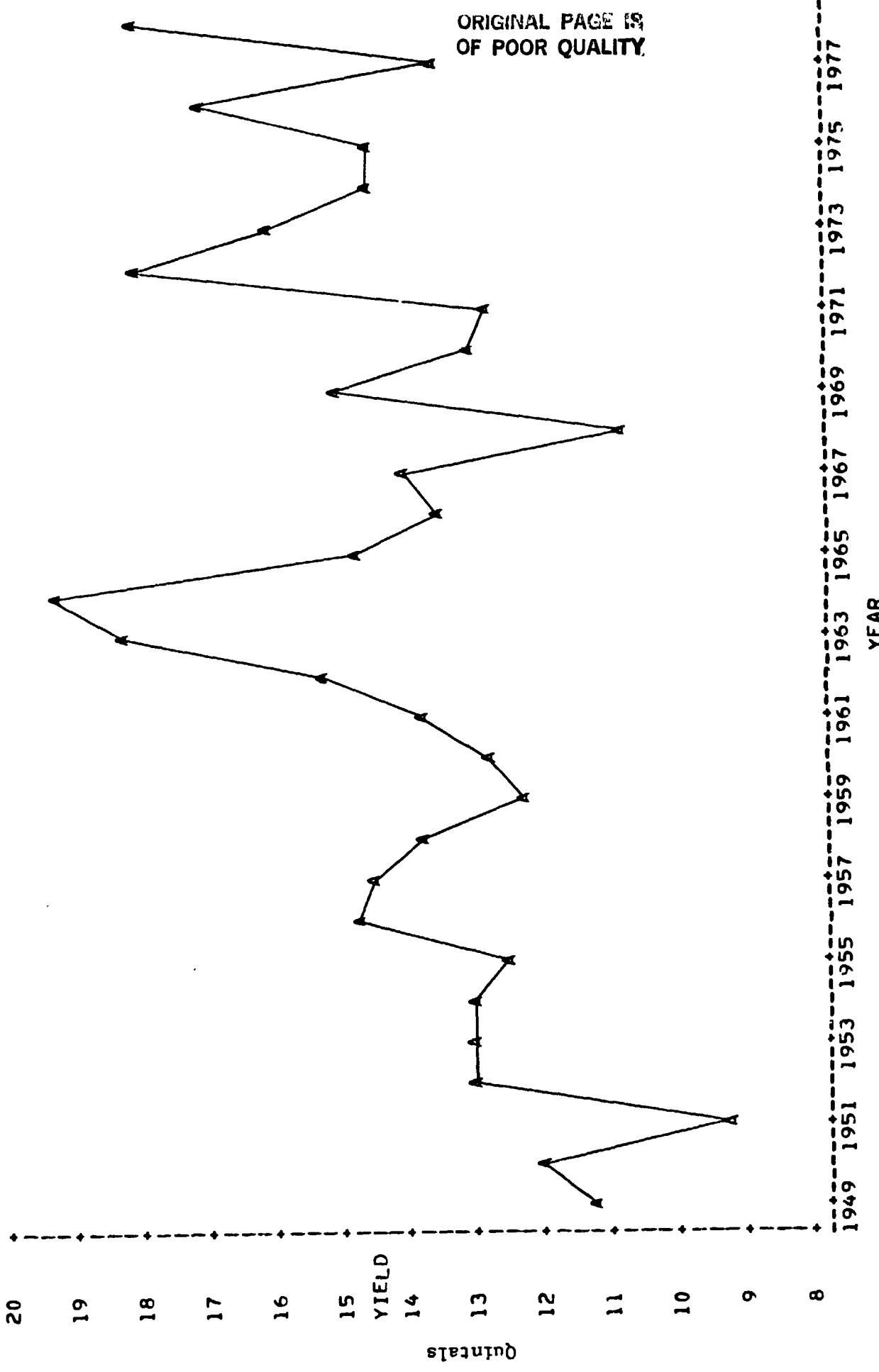


Figure 3. Plot of Yield Versus Year for Province of Buenos Aires.

effect on yield. A large, positive Z-index in August during jointing is favorable to yield. High temperatures in September during heading reduce yield. The statistics of the selected model are summarized in Table 2. A plot of the model's predicted yields and the actual yields for 1950-1970 is shown in Figure 4.

Cordoba

For Cordoba, plots of possible weather variables vs. yield showed that 1951 was a "pivot point" in determining linear relationships for most of the variables. Since yield for 1951 was extremely low, it was decided to plot the variables without 1951. No improvements was seen; there were even fewer linear relationships. Therefore, it was decided to retain 1951.

Initial variables were chosen on the basis of correlation with yield and were tried in regression equations. Models were tried with several combinations and squares of some of the variables. The following variables were significant as a model:

ETMETH5S.....	May ET minus \hat{ET} squared
TX7.....	July average maximum temperature
P-PET8.....	August precipitation minus PET
ETMETH90S.....	September-October average ET minus \hat{ET} squared.

The coefficient for ETMETH5S was negative, which reflects the need for drier conditions at planting time. The coefficient for ETMETH90S was also negative. This reflects the detrimental effects of excessive spring rains. High temperatures at early growth (July) reduces yield; favorable rain after planting (August) helps the crop get a good start. The statistics of the selected model are summarized in Table 3. A plot of the model's predicted yields and the actual yields for 1950-1970 is shown in Figure 5.

Entre Rios

Entre Rios was similar to Buenos Aires in that plots of weather variables versus yield did not produce any obvious linear relationships or strong

correlations. Regression equations containing moisture variables consistently produced negative coefficients for August through November. This can be explained agronomically by the fact that Entre Rios is a very humid province with a high annual rainfall. Rainfall greater than the normal expected value produces disease and fungus problems, thereby reducing yield.

Working with variables chosen on the basis of correlation with yield failed to produce satisfactory results. Therefore, a mechanical process using SAS procedures was used to narrow down possible variables. The best model obtained was:

ETMETH5.....May ET minus \hat{ET}
R6.....June total precipitation
ZINDEX8.....August Z index
R11.....November total precipitation.

All coefficients were negative for June, August and November, indicating the less rain the better. The statistics of the selected model are summarized in Table 4. A plot of the model's predicted yields and the actual yields for 1950-1970 is shown in Figure 6.

La Pampa

For La Pampa, a plot of yield vs. year showed 1965 yield to be extraordinarily high--so high as to be questionable. Correlations and plots of variables with yield were more favorable without 1965. Therefore, 1965 was eliminated from the model data set. The best model had the following variables:

ETMETH4.....April ET minus ET
TX8.....August average maximum temperature
ETMETH9.....September ET minus ET
P-PET10.....October precipitation minus PET
P-PET11.....November precipitation minus PET.

The problem with this and the other favorable models is that the coefficient for ETMETH4 was always negative. This cannot be explained agronomically. Therefore, the model was tried in the regression equation without ETMETH4. The R² was reduced, but all coefficients were reasonable. The

coefficient for the August variable was the only negative. The statistics of the selected model are summarized in Table 5. A plot of the model's predicted yields and the actual yields for 1950-1970 is shown in Figure 7.

Santa Fe

For Santa Fe, 1964 was a year of outstanding yield. Eliminating 1964 from the data set produced better plots and correlations of weather variables to yield.

Santa Fe was the only province for which a satisfactory model was not obtained. Several different approaches were taken to derive a reasonable model, but the only acceptable model was the following:

ETMETH5.....May ET minus \hat{ET}
ETMETH7.....July ET minus \hat{ET}
ETMETH10.....October ET minus \hat{ET}
ETMETH11.....November ET minus \hat{ET} .

In this model, ETMETH10 was significant only at the 20 percent level. It was decided to keep it in the model to reflect conditions at the critical heading period in October. In addition, the R^2 for the model without ETMETH10 was lower. The only negative coefficient was for the November variable. This reflects the need for drier conditions at harvest time. The statistics of the selected model are summarized in Table 6. A plot of the model's predicted yields and the actual yields for 1950-1970 is shown in Figure 8.

TEST RESULTS

A bootstrap test was run on each model. In this test, the last year of the yield data set was left out and the model was used to predict yield for that year. This process was repeated for the number of years desired. In this case, the number of years used in the test for each province depended on the number of years with complete data. For Santa Fe this was nine years, for the other four provinces six.

For Buenos Aires, Entre Ríos, and La Pampa the bootstrap test adequately predicted yields for 1971 through 1976. There were exceptions (a different

year in each case) where the difference between predicted and actual yield was rather large. For Cordoba, the model predicted yield accurately for three out of the six years and rather poorly for two of the years. In 1972 there was a nine quintal per hectare difference between actual and predicted yield with the predicted yield the lower of the two. In that particular year the beginning of the growing season was extremely dry with increased moisture late in the season, enabling a come-back for the wheat. For Santa Fe, although the model showed actual and predicted yields running fairly close for the modeling years, the bootstrap test failed to predict within five quintals seven of the eight test years. Bootstrap results and plots of model results are printed beginning on page 15.

APPENDIX

Definition of Variables

Three types of indices were used to measure amount of moisture available for plant growth. The first, P-PET, is a measure of precipitation minus potential evapotranspiration. Potential evapotranspiration is determined by the procedures developed by Thornthwaite (1948). It requires temperature only:

$$PET = \left(\frac{10T}{I} \right)^a$$

where I = heat index, which is the sum of the 12 monthly indices i ,

$$i = \left(\frac{T}{5} \right)^{1/5/4}$$

T = monthly temperature in $^{\circ}\text{C}$, and

$$a = \text{an empirical exponent} = 6.75 \times 10^{-7}I^3 - 7.71 \times 10^{-5}I^2 + 1.79 \times 10^{-2}I + 0.49.$$

The duration of daylight is used to adjust potential evapotranspiration as a portion of 12 hours.

The second index, Z index, is derived by an algorithm using monthly temperature and precipitation and is defined as

$$Z = dk$$

where $d = P - \hat{P}$ \hat{P} is the observed precipitation
 P is the climatically appropriate precipitation and is equal to $ET + R + R0 + L$.

Evapotranspiration \hat{ET} , recharge \hat{R} , runoff $\hat{R0}$, and loss \hat{L} are obtained by multiplying each of their respective potential values (PET, PR, PRO, PL) by the coefficient which is the ratio of their average values to their average potential values; that is, $\alpha = \bar{ET}/PET$, $\beta = \bar{R}/PR$, $\gamma = \bar{R0}/PRO$, and $\sigma = \bar{L}/PL$. Climatically appropriate evapotranspiration, recharge, runoff, and loss are then determined as $\hat{ET} = \alpha \cdot PET$; $\hat{R} = \beta \cdot PR$; $\hat{R0} = \gamma \cdot PRO$; and $\hat{L} = \sigma \cdot PL$. Recharge, runoff, and loss are determined by a hydrologic procedure developed by Palmer (1965).

The third index is the difference between ET and \hat{ET} . Soil moisture

depletion is based on evapotranspiration (ET) estimates, determined as follows:

$$(ET)_n = \frac{(S)_{n-1}}{AWC} \left[\{ (PET)_n - (P)_n \} + (P)_n \right]$$

where

$(ET)_n$ = "Actual" evapotranspiration,

$(S)_{n-1}$ = Available moisture at end of $n-1$ month,

AWC = Maximum water holding capacity,

$(P)_n$ = Precipitation for month n , and

$(PET)_n$ = Potential evapotranspiration for month n .

ET - ET measures the difference between the actual evapotranspiration and the "climatically appropriate" evapotranspiration, and hence gives an indication of soil moisture supply and demand.

Statistical Analysis System Procedures Used

PROC CORR	Computes correlation coefficients between variables, including Pearson product-moment and weighted product-moment correlation.
PROC PLOT	Graphs one variable against another, producing a printer plot.
PROC STEPWISE	Provides five methods for stepwise regression. Stepwise is useful when selecting variables to be included in a regression model from a collection of independent variables.
PROC STEPWISE FORWARD	Begins by finding the one-variable model that produces the highest R^2 . For each of the other independent variables, FORWARD calculates F-statistics reflecting the contribution to the model if the variables were to be included.
PROC STEPWISE BACKWARD	Begins by calculating statistics for a model including all the independent variables. The variables are deleted from the model one by one until all the remaining variables produce F-statistics significant at the .10 level.
PROC STEPWISE STEPWISE	The stepwise method is a modification of the forward selection technique, and differs in that variables already in the model do not necessarily stay there. After a variable is added (as in the forward selection method) the stepwise method looks at all the variables

already included in the model and deletes any variable that does not produce an F-statistic significant at the .10 level. Only after this check is made and the necessary deletions accomplished can another variable be added to the model.

PROC STEPWISE MAXR

(Maximum R^2 improvement) Unlike the three techniques above, this method does not settle on a single method. Instead it looks for the "best" two-variable model, the "best" three variable model, and so forth.

PROC PETM

Uses latitude and mean monthly temperatures to calculate Thornthwaite's potential evapo-transpiration for each month.

PROC ZINDEX

Uses monthly PET's, precipitation, SS (beginning moisture in surface layer), AWCS (available water capacity in surface layer), SU (beginning moisture in the underlying layer), and AWCU (available water capacity in the underlying layer) to calculate Palmer's soil moisture budget, drought index Z, ET, and ET.

STATISTICAL ANALYSIS SYSTEM 16:34 WEDNESDAY, NOV

BACKWARD ELIMINATION PROCEDURE FOR DEPENDENT VARIABLE YIELD

ALL VARIABLES ENTERED	R SQUARE = 0.87633601	C(P) = 0.0000000	F	PROB>F
REGRESSION	5	84.53539563	16.90706713	
ERROR	15	11.92918418	0.79527921	
TOTAL	20	96.46454381	21.26	0.0001
		STD ERROR	TYPE II SS	F
INTERCEPT	20.95565645	0.04894149	34.61436006	30.95
TREND	0.27255560	0.03967002	38.56080320	48.49
ETMETH4	0.37623307	0.0131750	19.60404949	0.0001
R7	-0.03932946	0.00680040	18.37554505	12.08
ZINDEX8	0.03268848	0.15058974	16.55851999	0.0002
TX9	-0.43245258		8.25	0.0116

LES IN THE MODEL ARE SIGNIFICANT AT THE 0.1000 LEVEL.

ORIGINAL VALUE OF
OF POOR QUALITY

Table 2. Statistics of Model for Province of Buenos Aires.

STATISTICAL ANALYSIS SYSTEM

16:40 WEDNESDAY NOV.

BACKWARD ELIMINATION PROCEDURE FOR DEPENDENT VARIABLE YIELD

ALL VARIABLES ENTERED	R SQUARE = 0.69487993	C (P) = 5.0000000	MEAN SQUARE	F	PROB>F
DF	SUM OF SQUARES				
REGRESSION	4	134.31095736	33.57774184	9.11	0.0005
ERROR	16	158.99508978	33.68719311		
TOTAL	20	193.30605714			
	B VALUE	STD ERROR	TYPE II SS	F	PROB>F
INTERCEPT	21.03908139	0.01986689	13.66259933	3.78	0.0887
FTYETHSS	-0.02395604	0.02914134	40.91690661	11.10	0.0042
XPFT8	0.09700964	0.0294987	40.93871412	11.10	0.0042
ETMET90S	-0.00982928	0.00294987			

LES IN THE MODEL ARE SIGNIFICANT AT THE 0.1000 LEVEL.

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Table 3. Statistics of Model for Province of Cordoba.

BUENOS AIRES ACTUAL YIELD AND MODEL PREDICTION 1950-1970
A=ACTUAL YIELD
P=MODEL YIELD

16:39 WEDNESDAY, NOVEMBER 11, 1970

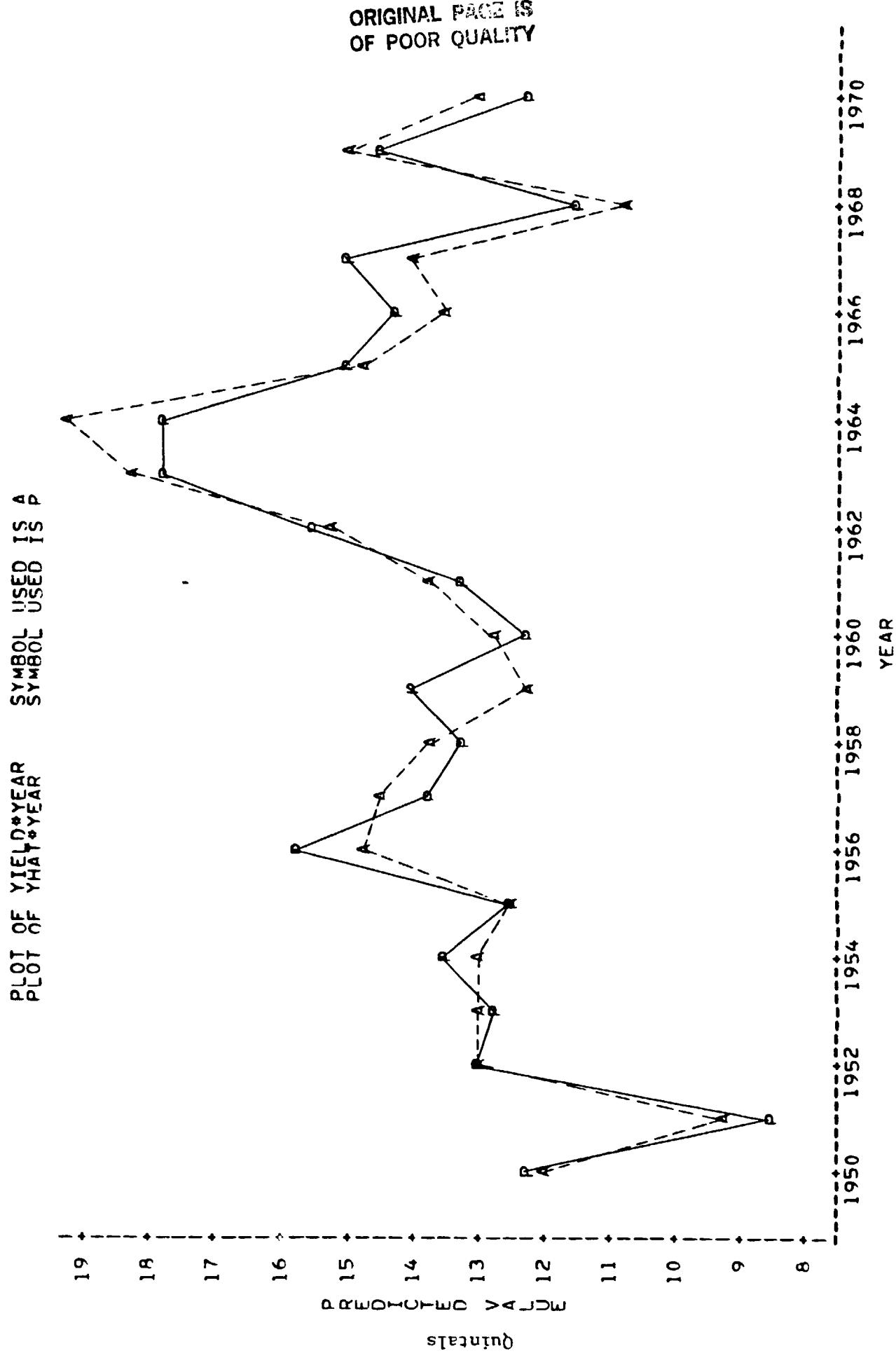


Figure 4. Province of Buenos Aires Wheat Model.

CORDOBA ACTUAL YIELD AND MODEL PREDICTION 1950-1970

16:40 WEDNESDAY, NOVEMBER 11, 1970

$A = \text{ACTUAL YIELD}$
 $P = \text{MODEL YIELD}$

PLLOT OF YIELD*YEAR
 PLLOT OF $\hat{Y}_t^*Y_t^*$ YEAR

ORIGINAL PLOTS
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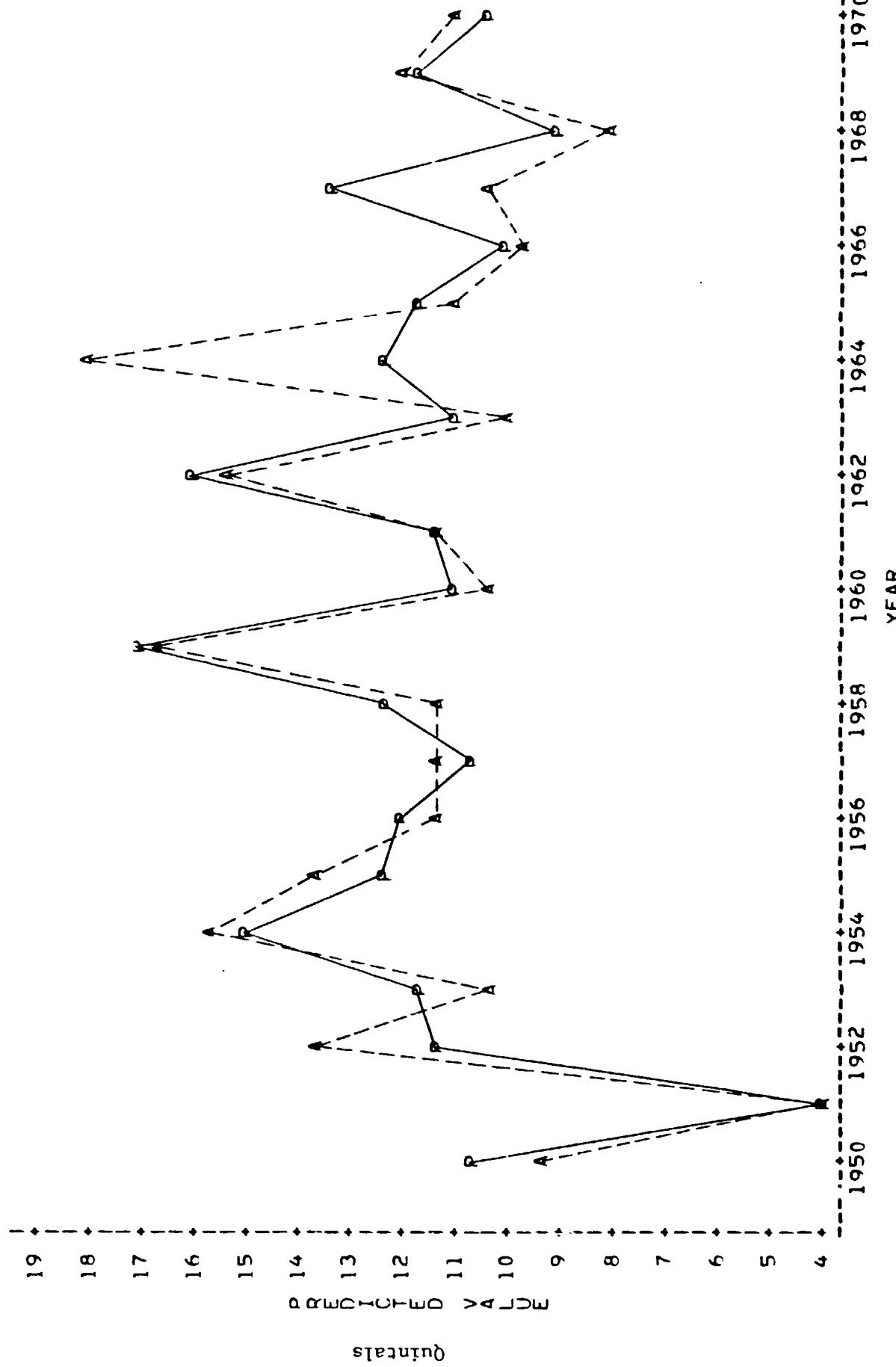


Figure 5. Province of Cordoba Wheat Model.

STATISTICAL ANALYSIS SYSTEM

DATA FOR PREDICTION PURPOSES FOR DIFFERENT VARIETIES

ALL VARIETIES	STN	YIELD	STN	YIELD
0.55636504	14	56.67871962	13.91967415	5.00000000
0.55636504	21	42.64037462	17.66563616	5.00000000
0.55636504	11	95.31424524	11.5157115	5.00000000
0.55636504	17	11.5157115	0.00000000	5.00000000
0.55636504	26	0.00000000	15.10461351	5.00000000
0.55636504	27	0.00000000	21.47461351	5.00000000
0.55636504	13	0.00000000	10.51646174	5.00000000
0.55636504	10	0.00000000	12.81570570	5.00000000

LEAST SIGNIFICANT DIFFERENCE AT THE 0.1000 LEVEL.

LEAST SIGNIFICANT DIFFERENCE AT THE 0.05 LEVEL.

LEAST SIGNIFICANT DIFFERENCE AT THE 0.01 LEVEL.

LEAST SIGNIFICANT DIFFERENCE AT THE 0.001 LEVEL.

LEAST SIGNIFICANT DIFFERENCE AT THE 0.0001 LEVEL.

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Table 4. Statistics of Model for Province of Batang Ries.

FIG. 5. PLOTS ACTUAL YIELD AND MULTI. PRODUCTION 1950-1970
A=ACTUAL YIELD
B=MULTI. PROD.

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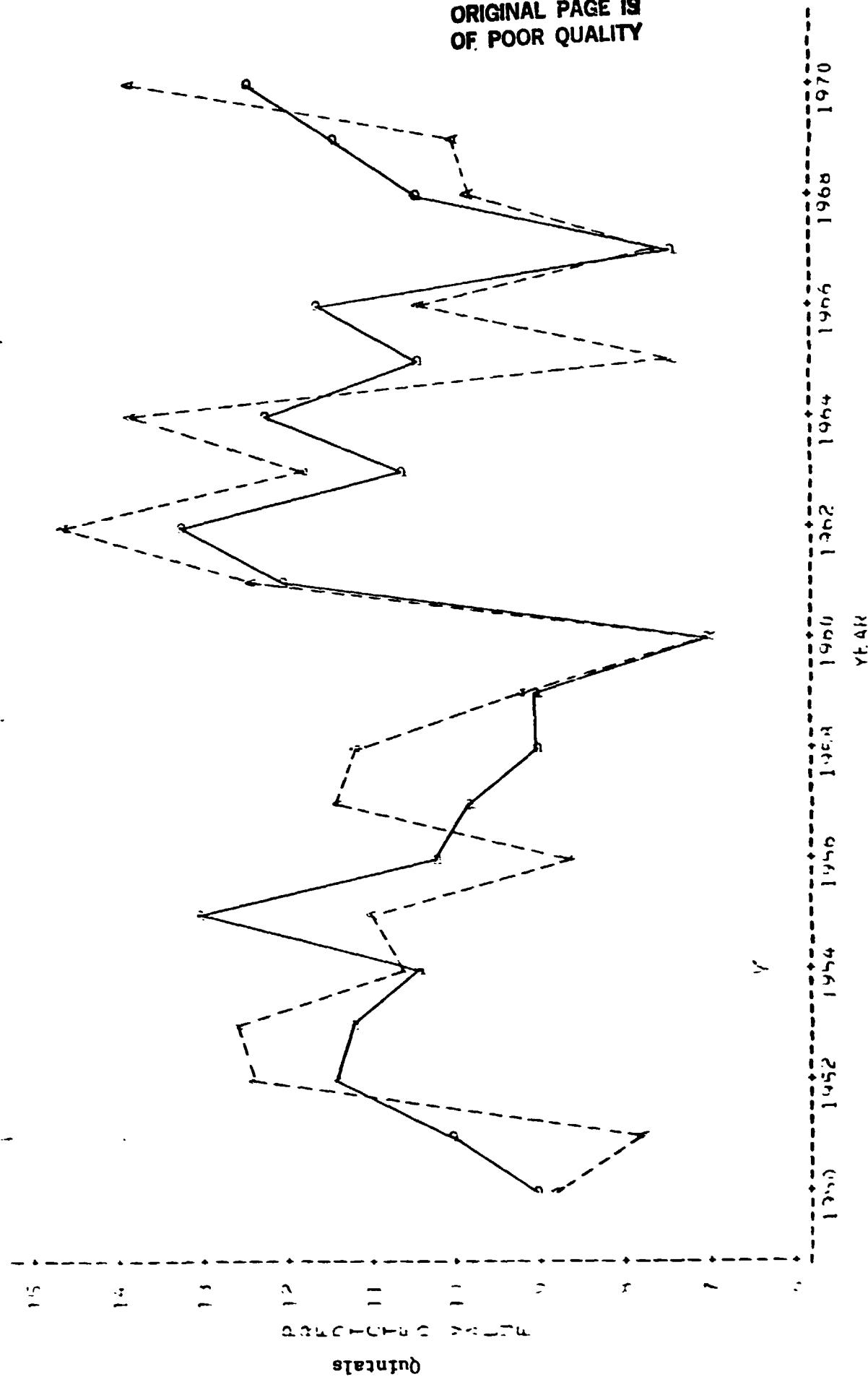


Figure 6. Province of Entre Ríos Wheat Model.

STATISTICAL ANALYSIS SYSTEM

15:15 WEDNESDAY.

BACKWARD ELIMINATION PROCEDURE FOR DEPENDENT VARIABLE YIELD	
ALI. VARIABLE ENTERED	D. SQUARE = 0.7306214
INTERCEPT	C(P) = 5.0000000
TXR	DF
FT'FETHQ	SUM OF SQUARES
B_PET10	117.55938273
P_PFT11	43.2907227
	160.8501750
	MEAN SQUARE
	29.38992568
	2.8860515
	PROB>F
	1n.1H 0.0001
	TYPE II SS
	STN ERROR
	H VALUE
INTERCEPT	31.44093043
TXR	-1.24177470
FT'FETHQ	0.34140295
B_PET10	0.02426685
P_PFT11	0.03423129

P VALUES IN THE MODEL ARE SIGNIFICANT AT THE 0.1000 LEVEL.

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Table 5. Statistics of Model for Province of La Pampa.

1. A PAMPA ACTUAL YIELD AND MODEL PREDICTION 1950-1970
A=ACTUAL YIELD
P=MODELS YIELD

PLOT OF YIELD*YEAR SYMBOL USED IS A

15:15 WEDNESDAY.

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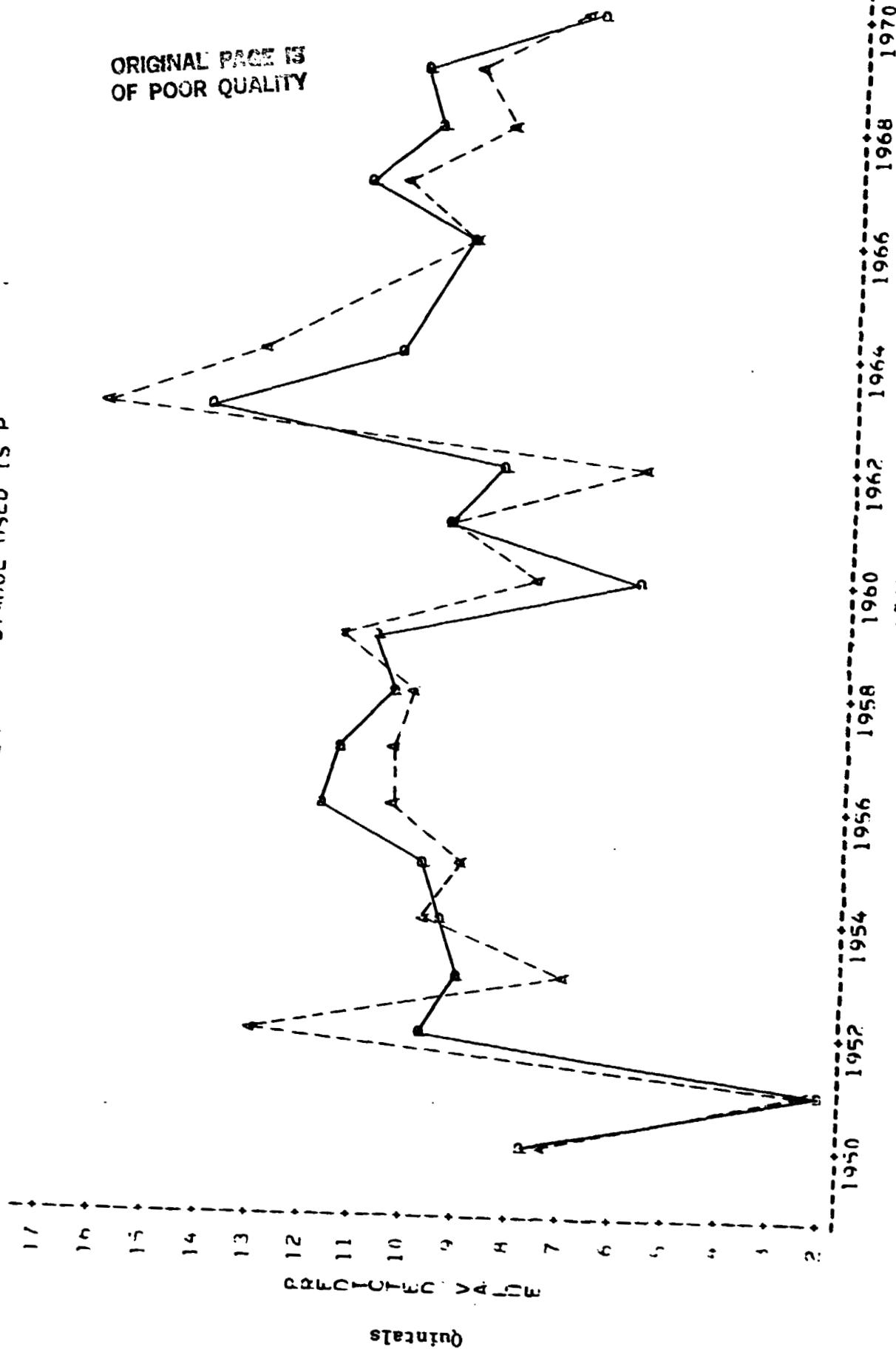


Figure 7. Province of La Pampa Wheat Model.

STATISTICAL ANALYSIS SYSTEM

16:42 WEDNESDAY, NOV

BACKWARD ELIMINATION PROCEDURE FOR DEPENDENT VARIABLE YIELD

ALL VARIABLES ENTERED	R SQUARE = 0.73525485	C(P) = 5.0000000	PROB>F	
	DF	SUM OF SQUARES	MEAN SQUARE	F
REGRESSION	4	121.96620241	30.49156560	10.41
ERROR	15	43.91671259	2.9277804	0.0003
TOTAL	19	165.88291500		
	B VALUE	STD ERROR	TYPE II SS	F
INTERCEPT	14.21633800	0.07842409	16.51007703	5.64
ETMETH5	0.18631292	0.32315203	0.18399707	0.0313
ETMETH7	1.58218278	0.12021388	5.22936706	0.0002
ETMETH10	0.16044098	0.22924733	17.23496053	0.2013
ETMETH11	-0.55618748		5.89	0.0283

VARIABLE ETMETH10 REMOVED R SQUARE = 0.70373042 C(P) = 4.78611971

	DF	SUM OF SQUARES	MEAN SQUARE	F	PROB>F
	B VALUE	STD ERROR	TYPE II SS	F	PROB>F
REGRESSION	3	116.73684535	38.91229845	12.67	0.0002
ERROR	16	49.14607465	3.07162998		
TOTAL	19	165.88291500			
INTERCEPT	14.21688607	0.07915607	13.87136650	4.52	0.0495
ETMETH5	0.16521294	0.33081439	69.20214824	22.53	0.0002
ETMETH7	1.57050198	0.22634741	24.34871304	7.93	0.0124
ETMETH11	-0.63741900				

LES IN THE MODEL ARE SIGNIFICANT AT THE 0.1000 LEVEL.

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Table 6. Statistics of Model for Province of Santa Fe.

SANTA FE ACTUAL YIELD AND MODEL PREDICTION 1950-1970

A=ACTUAL YIELD
P=MODELS YIELD

PLOT OF YIELD*YEAR SYMBOL USED IS A
PLOT OF YHAT*YEAR SYMBOL USED IS P

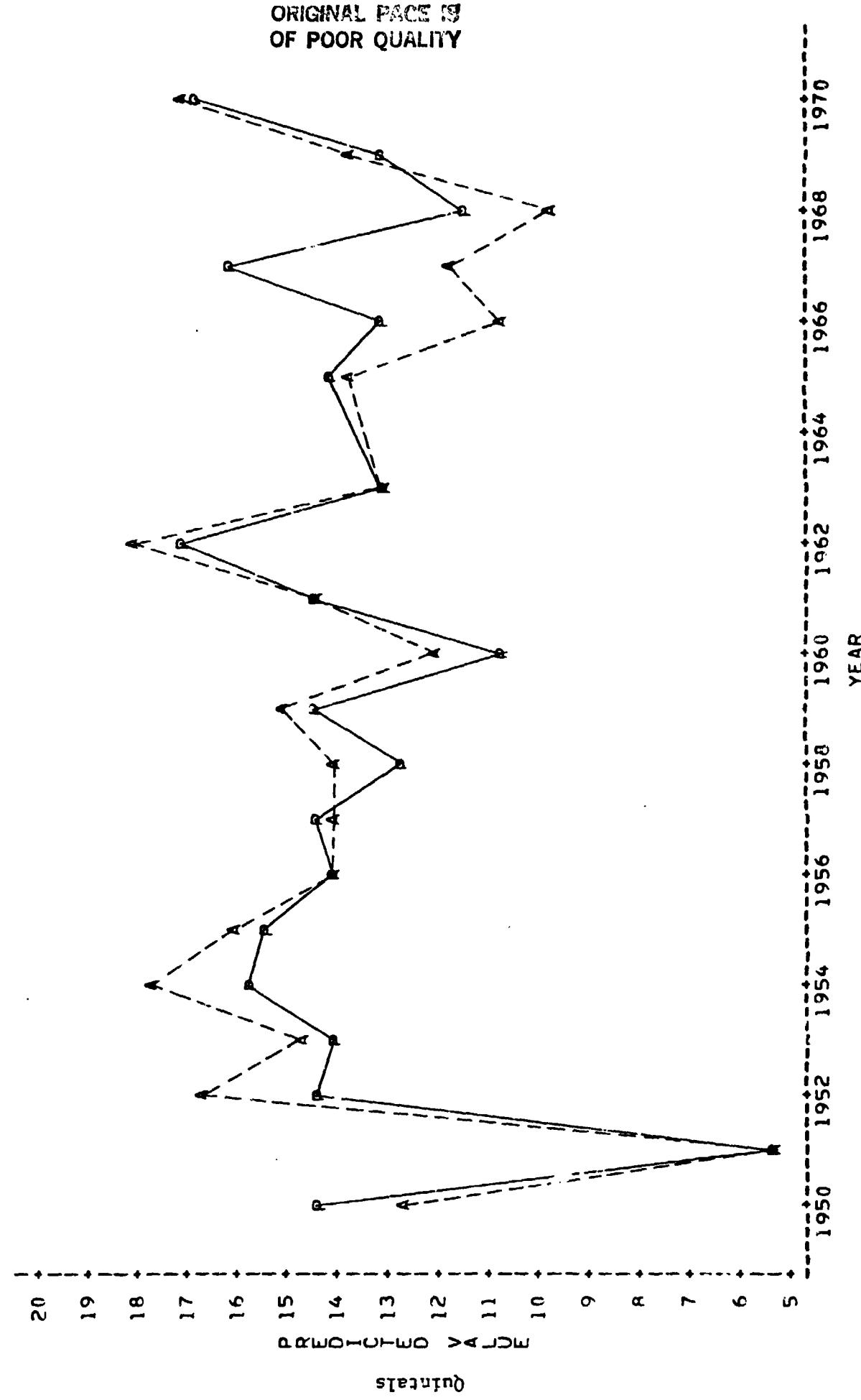


Figure 8. Province of Santa Fe Wheat Model.

STATISTICAL ANALYSIS SYSTEM										15:54 MONDAY, JANUARY	
BETA2	HEGMIN YR	ENDMID YR	YIELD YR	YIELD YR	ORSYIELD	RSQ	PDER	DFRES	MSHES	HETA1	
0.273318	1960	1959	1970	11.9289	13.09	0.882743	0.541325	14	0.811310	22.2659	
0.272556	1960	1970	1971	17.3248	12.71	0.876136	0.35325	15	0.74528	22.09557	
0.260200	1959	1971	1972	13.8520	17.94	0.729944	0.33872	16	1.64340	22.1111	
0.294499	1959	1972	1973	14.4450	16.08	0.41539	0.649591	17	2.36177	22.34403	
0.314440	1959	1973	1974	12.4539	14.50	0.644128	0.629124	18	2.35986	22.35306	
0.316092	1959	1974	1975	15.4460	14.52	0.629124	0.27583	19	2.33339	22.42570	
0.326449	1959	1975	1976	16.8701	17.01	0.624703	0.94960	20	2.24949	24.4320	
HF1A3											
0.277284	-0.034817	0.0345076	-0.50808	22.2654	3.55313	0.0183	-0.7763	-2.0056	-11.127		
0.276233	-0.039353	0.0326485	-0.43245	20.9557	3.54322	0.6268	-1.4552	-2.0871	-18.433		
0.244091	-0.029661	0.014751	-0.53698	22.4011	3.38026	-1.3892	-0.1627	-0.1027	-10.073		
0.240360	-0.034419	0.0214974	-0.58838	23.4880	3.88035	-0.5547	-1.4599	-0.9567	-11.061		
0.210257	-0.036872	0.0177451	-0.59889	23.5306	4.08771	-2.3520	-2.1754	-0.0041	-10.241		
0.163264	-0.032536	0.0162190	-0.65615	24.2570	4.36920	-0.5640	-1.4785	-0.0052	-12.861		
0.179219	-0.031342	0.0157948	-0.60829	24.4320	4.32184	-0.4772	-1.0970	-1.3668	-12.631		

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Table 7. Results of Jackknife Test for Buenos Aires Wheat Model.

BUENOS AIRES BOOTSTRAP RESULTS--TEST OF MODEL 1970-1976
 O=OBSERVED YIELD
 P=MODELS PREDICTED YIELD
 PLOT OF YIELD vs. YIELD vs. SYMBOL USED IS P

16:39 WEDNESDAY, NOVEMBER

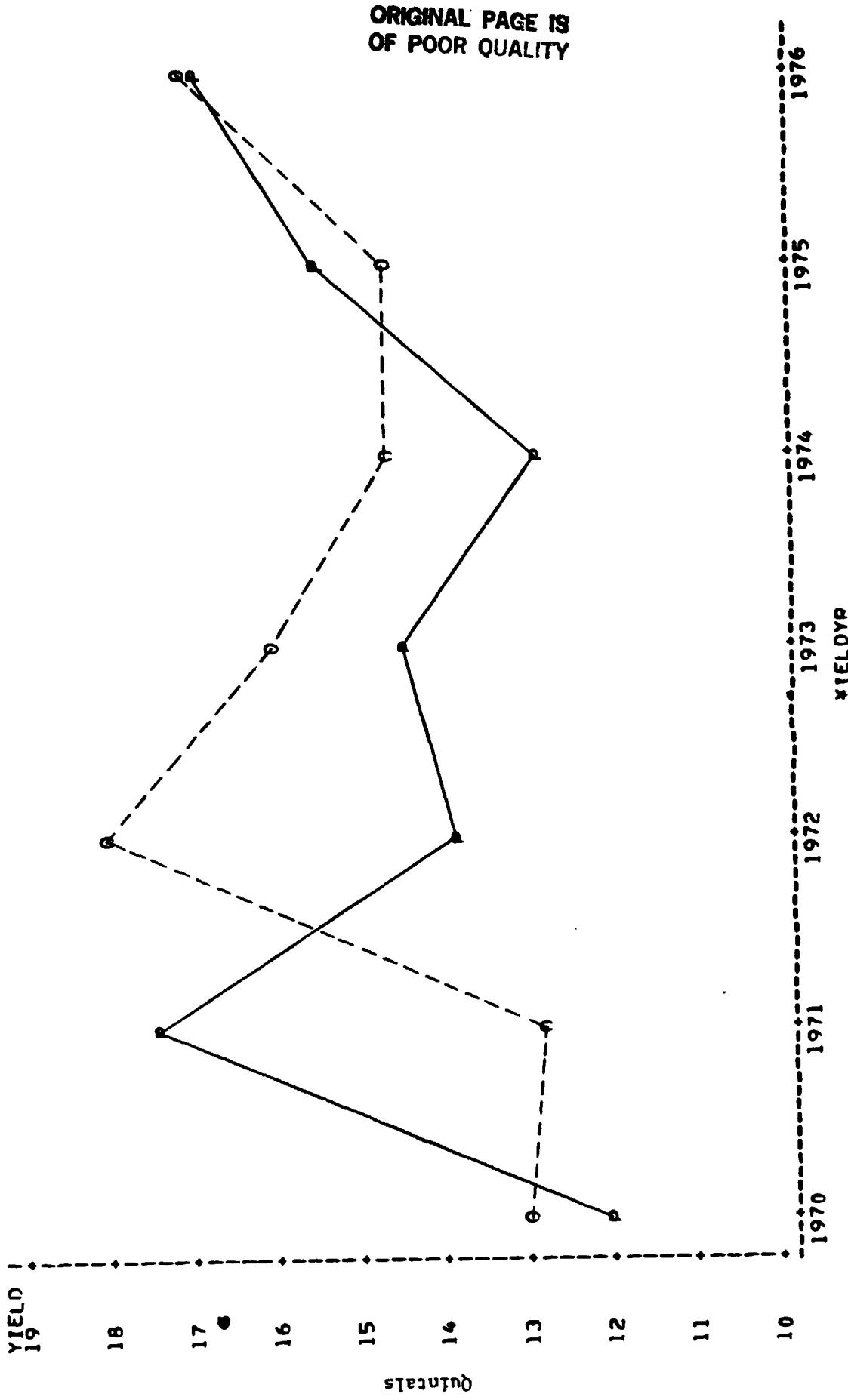


Figure 9. Province of Buenos Aires Wheat Model.

STATISTICAL ANALYSIS SYSTEM										16:00 MONDAY, JANUARY	
BETA1	BFGMDL.YR	ENMDLYR	YFLDLYR	YIELD	OBSYIELD	R50	POER	UFRES	MSRES	RUNI	
21.3036	1950	1969	1970	10.2936	10.97	0.66	0.8217	15	3.98781		
22.0.3375	1950	1970	1971	11.2243	12.66	0.696043	0.8623	16	33.657238		
22.0.7486	1950	1971	1972	13.0911	12.46	0.694816	0.3341	17	33.87298		
22.4.5507	1950	1972	1973	12.7246	12.25	0.647466	0.2333	18	32.439566		
22.4.6416	1950	1973	1974	12.8760	14.32	0.642316	0.216	19	32.757601		
22.6.0842	1950	1974	1975	13.0393	14.76	0.586285	0.4946	20	31.4951		
				13.6613	18.70	0.624300	0.951	21	32.961		
HFTA1	HFTA1	HFTA4	HFTA5	CONTR181	CONTR182	CONTR183	CONTR184	CONTR185	CONTR186	RUNI	
-0.017972	-0.44128	0.099569	-0.009637	21.3036	-0.171	-7.447	-2.8389	-0.1302			
-0.019428	-0.42396	0.097010	-0.009429	21.0371	-7.0210	-8.122	-1.4126	-0.0368			
-0.019393	-0.47531	0.096163	-0.010234	20.9317	-11.446	-6.179	-0.5726	-0.245466			
-0.004613	-0.42195	0.098008	-0.009615	20.7407	-10.062	-5.317	-0.5947	-0.04697			
-0.004717	-0.64690	0.083075	-0.009347	24.5507	-0.098	-10.544	-0.5408	-0.0937			
-0.004958	-0.64743	0.093695	-0.009465	24.6416	-0.068	-9.913	-0.4758	-0.0904			
-0.006079	-0.70276	0.109747	-0.009471	26.0842	-0.089	-11.455	-4.3213	-0.22046			

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Table 8. Results of Jackknife Test for Cordoba Wheat Model.

CORDORA BOOTSTRAP RESULTS--TEST OF MODEL 1970-1976
 O=OBSERVED YIELD
 P=MODEL PREDICTED YIELD
 PLOT OF YIELD*YIELD*YIELD*YIELD*YIELD*YIELD
 PLOT OF OBSERVED YIELD*YIELD*YIELD*YIELD*YIELD*YIELD

16:40 WEDNESDAY, NOVEMBER

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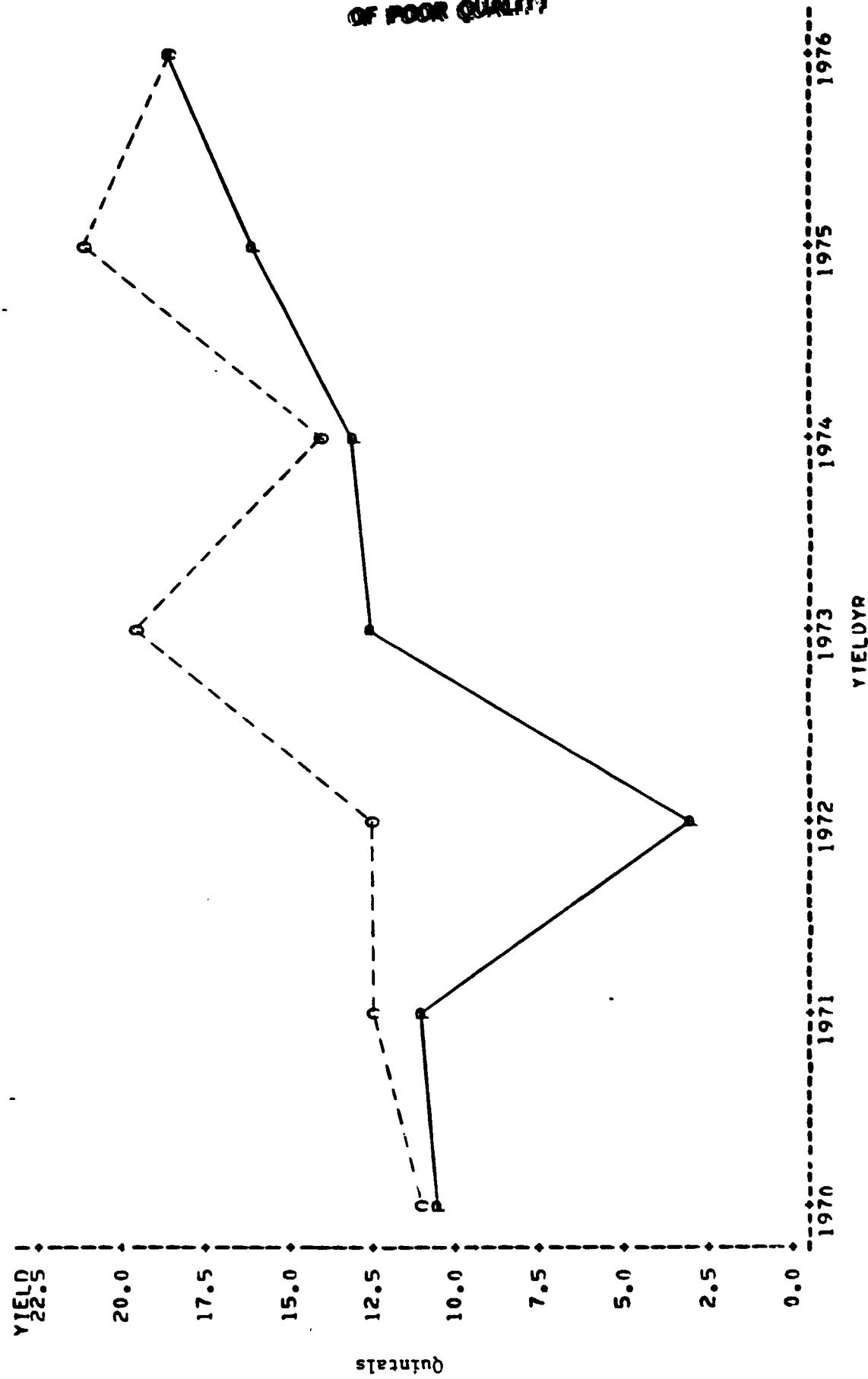


Figure 10. Province of Cordoba Wheat Model.

STATISTICAL ANALYSIS									
MONDAY, JANUARY		16:00		SYSTEM		POER		OFRES	
BETA1	MONDAY	YIELD	YIELD	OSYIELD	OSYIELD	OSYIELD	OSYIELD	OSYIELD	OSYIELD
13.1074	1969	1969	1970	13.88	13.88	13.88	13.88	13.88	13.88
13.0920	1969	1970	1971	10.28	10.28	10.28	10.28	10.28	10.28
13.5213	1970	1971	1972	9.75	9.75	9.75	9.75	9.75	9.75
13.5087	1970	1971	1971	11.95	11.95	11.95	11.95	11.95	11.95
13.4541	1970	1971	1974	16.11	16.11	16.11	16.11	16.11	16.11
13.4674	1970	1971	1975	11.39	11.39	11.39	11.39	11.39	11.39
8t.14.2	4t.14.3	8t.14.4	8t.14.5	10.67	10.67	10.67	10.67	10.67	10.67
0.026697	-0.021561	-0.017539	-0.012625	13.10/4	13.10/4	13.10/4	13.10/4	13.10/4	13.10/4
0.025534.9	-0.021569	-0.016837	-0.015834	13.61/4	13.61/4	13.61/4	13.61/4	13.61/4	13.61/4
0.023348.6	-0.022533	-0.016823	-0.015827	13.56/4	13.56/4	13.56/4	13.56/4	13.56/4	13.56/4
0.023417.6	-0.022330	-0.016494	-0.015358	13.59/4	13.59/4	13.59/4	13.59/4	13.59/4	13.59/4
0.027186.8	-0.021772	-0.022657	-0.015845	13.88/4	13.88/4	13.88/4	13.88/4	13.88/4	13.88/4
0.027220.7	-0.0217736	-0.022621	-0.014654	13.40/4	13.40/4	13.40/4	13.40/4	13.40/4	13.40/4

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Table 9. Results of Jackknife Test for Entre Ríos Wheat Model.

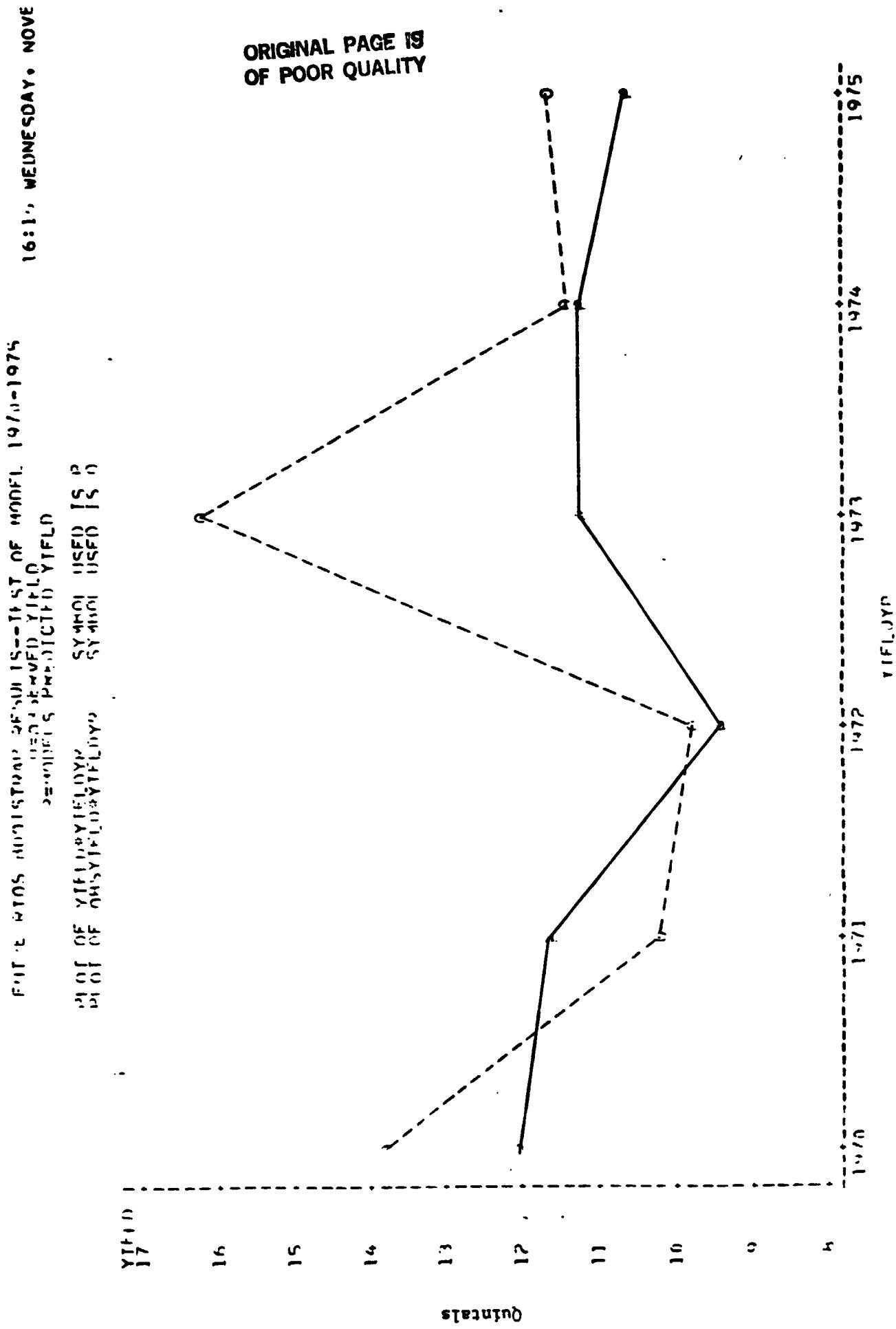


Figure 11. Province of Entre Ríos Wheat Model.

STATISTICAL ANALYSIS SYSTEM 15:35 WEDNESDAY JAN

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Table 10. Results of Jackknife Test for La Pampa Wheat Model.

LA PAMPA HONTSTRAP RESULTS--TEST OF MODEL 1970-1976
 P=0.05 PREDICTED YIELD
 PL.OT OF YIELD*YIELD*YIELD*YIELD*YIELD*YIELD

15:15 WEDNESDAY, J.

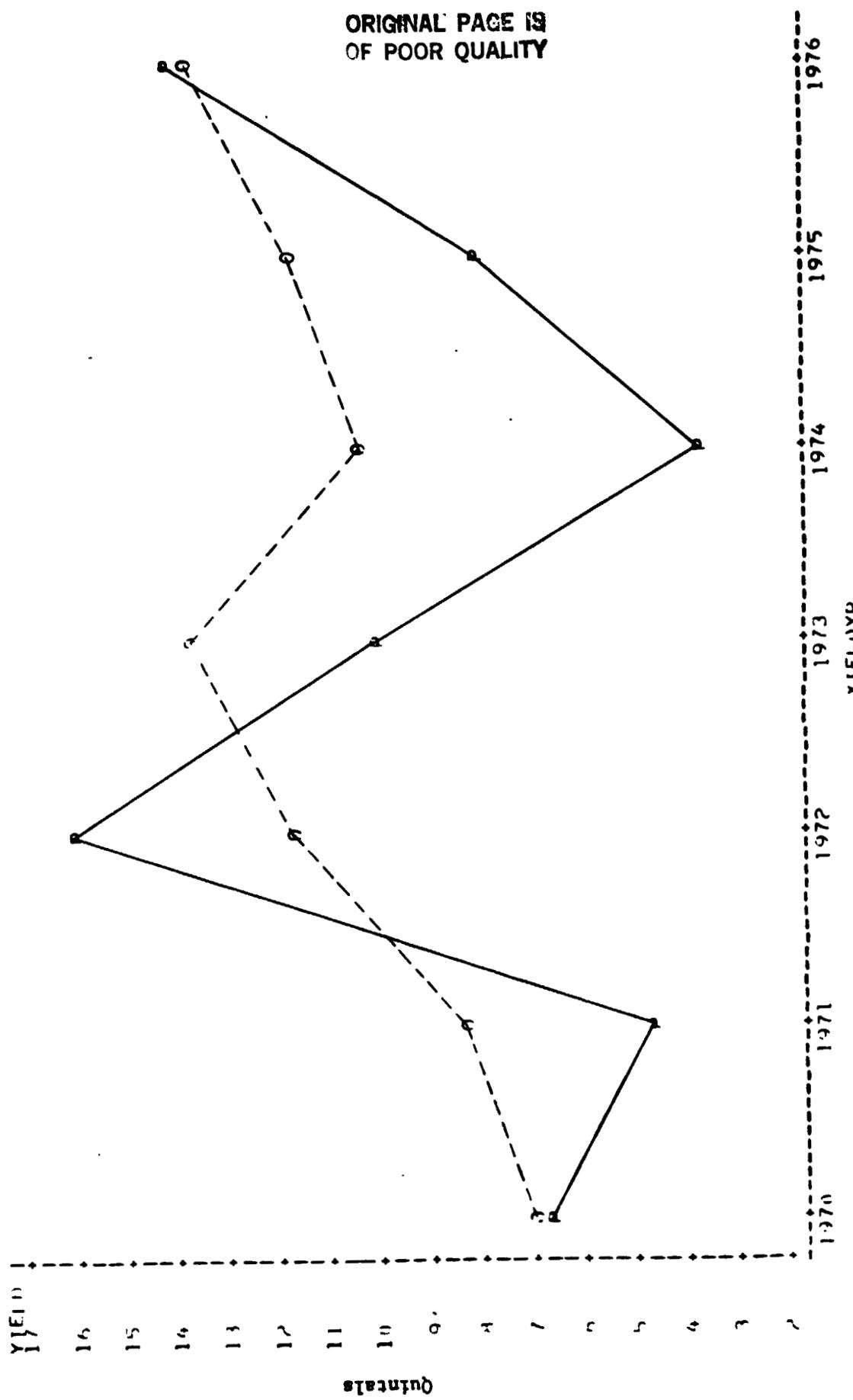


Figure 12. Province of La Pampa Wheat Model.

STATISTICAL ANALYSIS SYSTEM										16:01 MONDAY, JANUARY	
BETA1	BETA2	BETA3	BETA4	MEANLYR	YIELD	YIELD	NOSEYIELD	R50	POER	IPRES	MSRES
1.401592	1.950	1.960	1.970	1.604210	1.604210	1.604210	0.716705	0.8974	1.4	3.11587	
1.40163	1.950	1.971	1.971	1.600666	1.600666	1.600666	0.735255	0.8974	1.5	3.21577	
1.40162	1.950	1.971	1.972	1.02094	1.02094	1.02094	0.730648	2.4984	1.6	3.22177	
1.40134	1.950	1.972	1.974	1.41556	1.41556	1.41556	0.651933	3.4797	1.7	3.34797	
1.404530	1.950	1.974	1.974	2.305585	2.305585	2.305585	0.581830	2.807969	1.8	3.46126	
1.401349	1.950	1.974	1.975	1.55278	1.55278	1.55278	0.554184	2.062743	1.9	3.46126	
1.404251	1.950	1.975	1.975	1.52777	1.52777	1.52777	0.497853	0.5433	2.0	6.97634	
1.405165	1.950	1.976	1.977	1.09645	1.09645	1.09645	0.475850	1.8411	2.1	7.70584	
1.40090	1.950	1.977	1.978	1.52686	1.52686	1.52686	0.403816	0.5916	2.2	8.44353	

SANTA FE BOOTSTRAP RESULTS--TEST OF MODEL 1970-1978
 O=OBSERVED YIELD
 P=MODELS PREDICTED YIELD
 PLOT OF YIELD*YIELD*YR SYMBOL USED IS P
 PLOT OF YIELD*YIELD*YR SYMBOL USED IS O

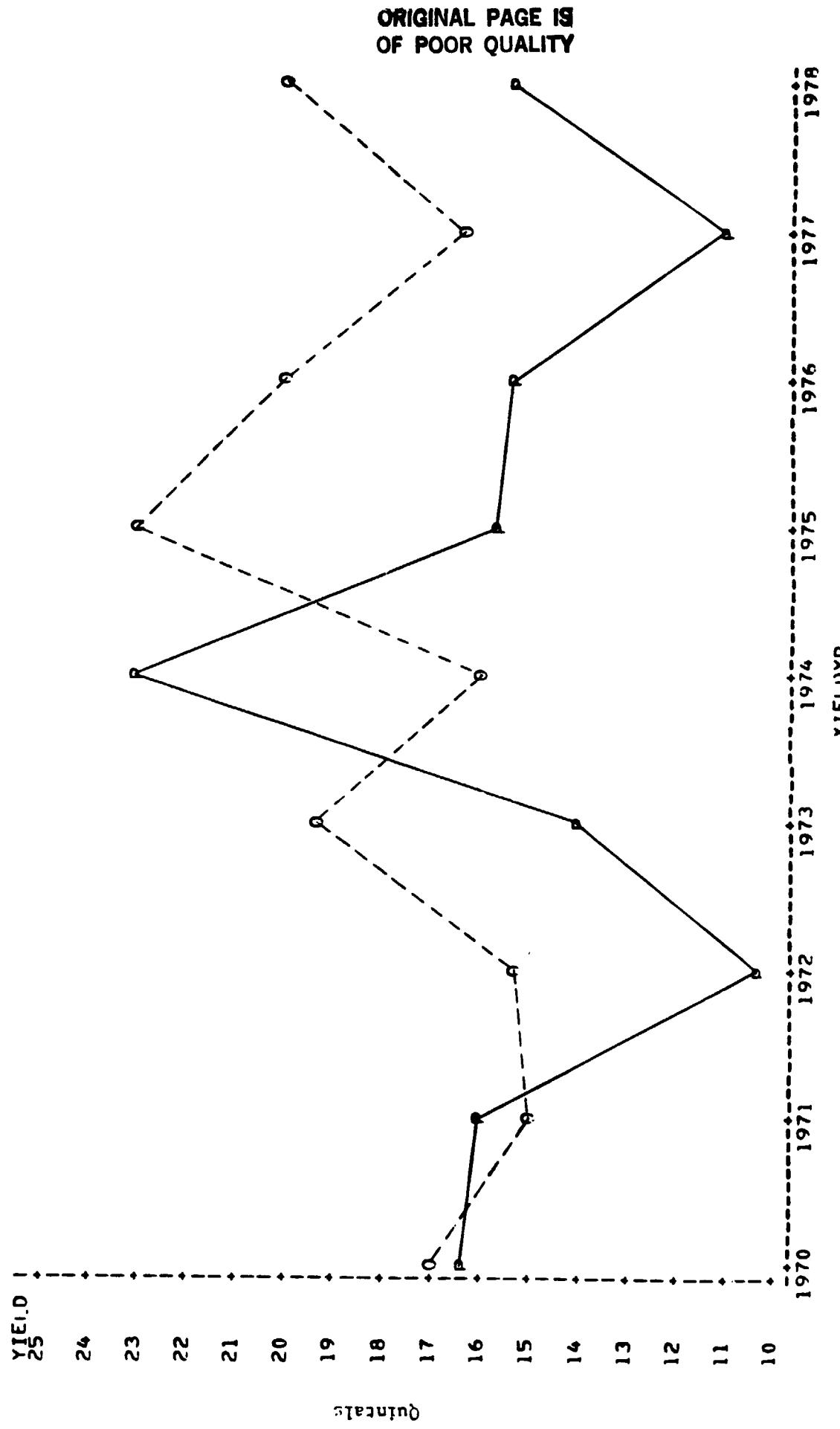


Figure 13. Province of Santa Fe Wheat Model.